

# Estimation of Moran's I in the Context of Uncertain Mobile Sensor Measurements

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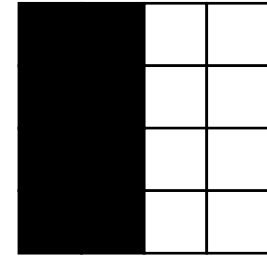
# Moran's I

$$I = \frac{n}{\sum_{i,j \neq i} w_{ij}} \cdot \frac{\sum_{i,j \neq i} w_{ij} (x_i - \bar{x})(x_j - \bar{x})}{\sum_i (x_i - \bar{x})^2}$$

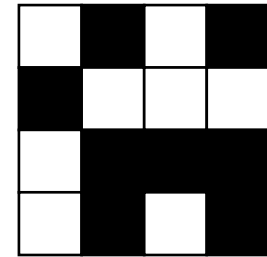
$x_i$ : observation at location i

$\bar{x}$ : mean of observations

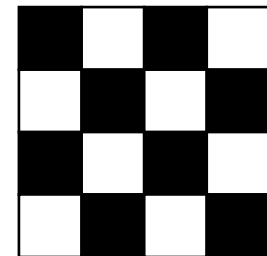
$w_{ij}$ : spatial weight between location i and j



I=1



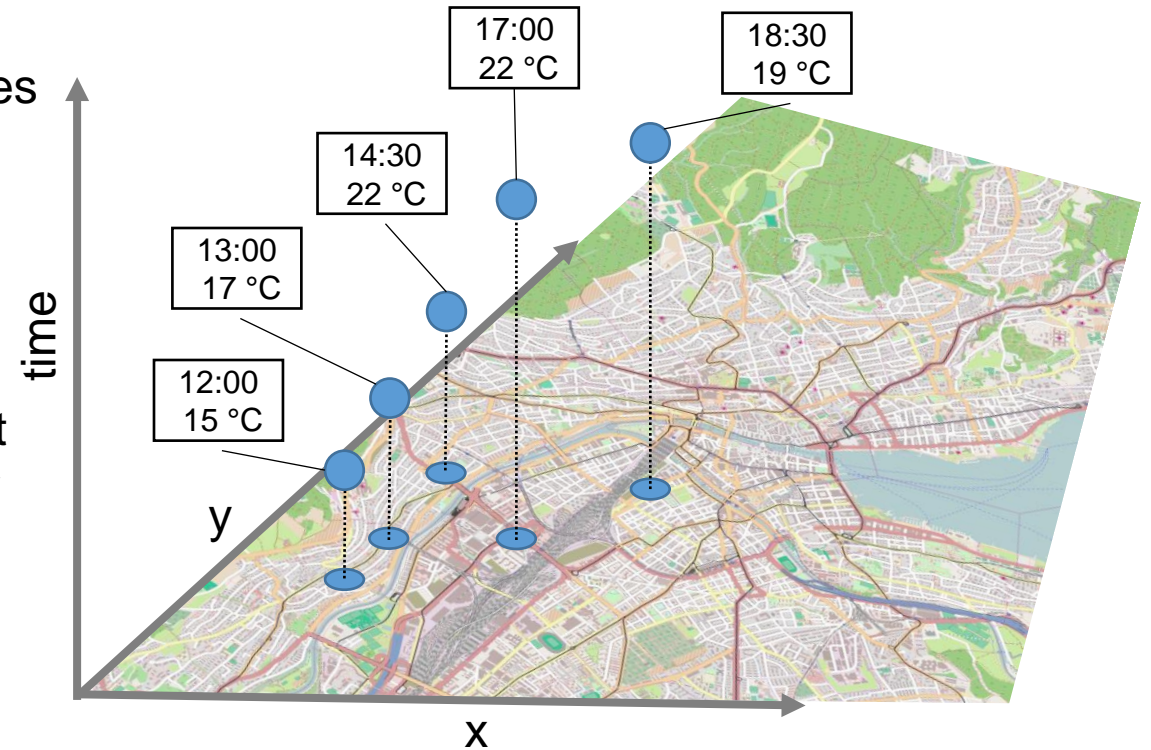
I=0



I=-1

# Motivation

- Mobile sensors capture measurements at different locations at different points in time and space.
- Measurements may represent different processes of dynamic geographic phenomena.
- The validity of measurements from the past at different locations is therefore uncertain.
- **Research question:** How can we calculate Moran's  $I$  with data collected at different times at different locations while taking into account their pair-wise joint uncertainty?



# Moran's $I$ with Modified Weights

Given a pairwise measure of certainty between the latest observations at two locations  $u_{ij}$  we define two sets of new weights which lead to modified Moran's  $I$  values

$$w'_{ij} = w_{ij} \cdot u_{ij} \quad \rightarrow I^1$$

$$w''_{ij} = w_{ij} \cdot (1 + u_{ij} - \bar{u}) \quad \rightarrow I^2$$

Intuition:

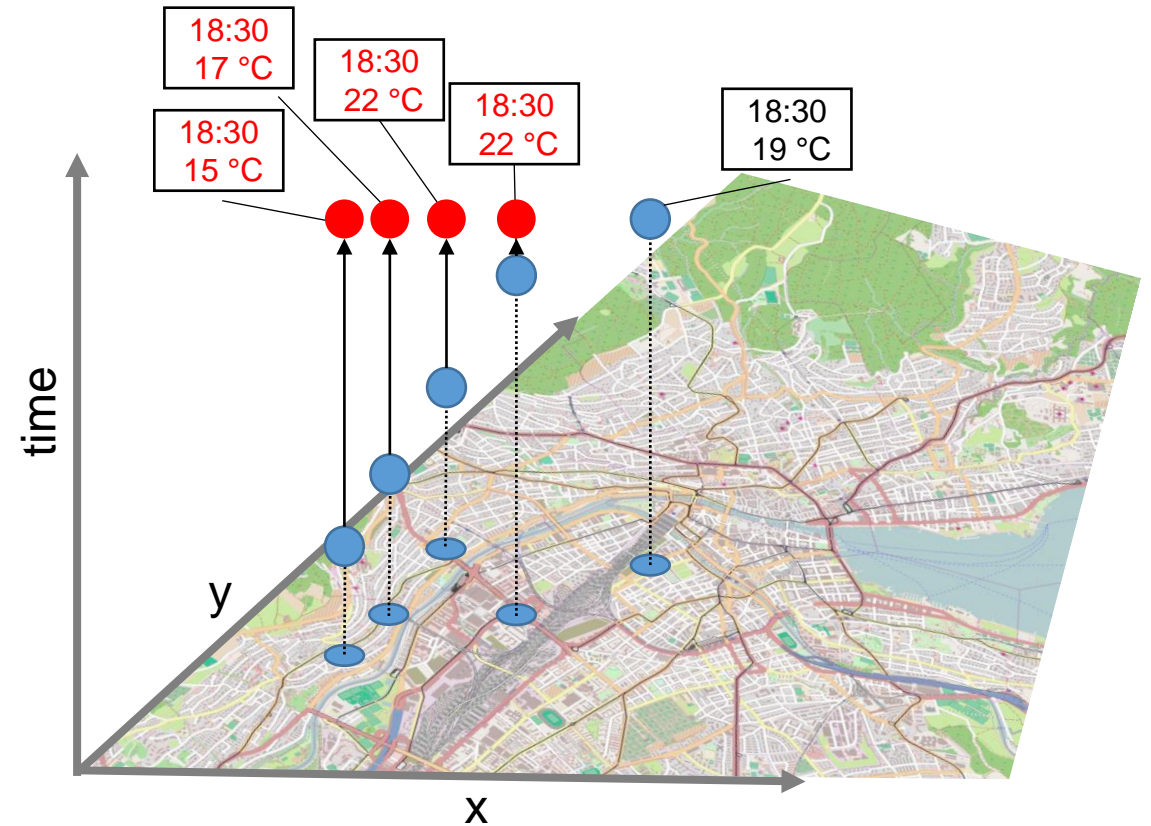
- $I^1$ : Absolute approach. Lower the weights of locations with observations with high certainty.
- $I^2$ : Relative approach. Locations with observations with above average certainty have increased weights, those with below average have reduced weights.

## Global Moran's $I$

$$I = \frac{n}{\sum_{i,j \neq i} w_{ij}} \cdot \frac{\sum_{i,j \neq i} w_{ij} (x_i - \bar{x})(x_j - \bar{x})}{\sum_i (x_i - \bar{x})^2}$$

# Uncertainty Estimation

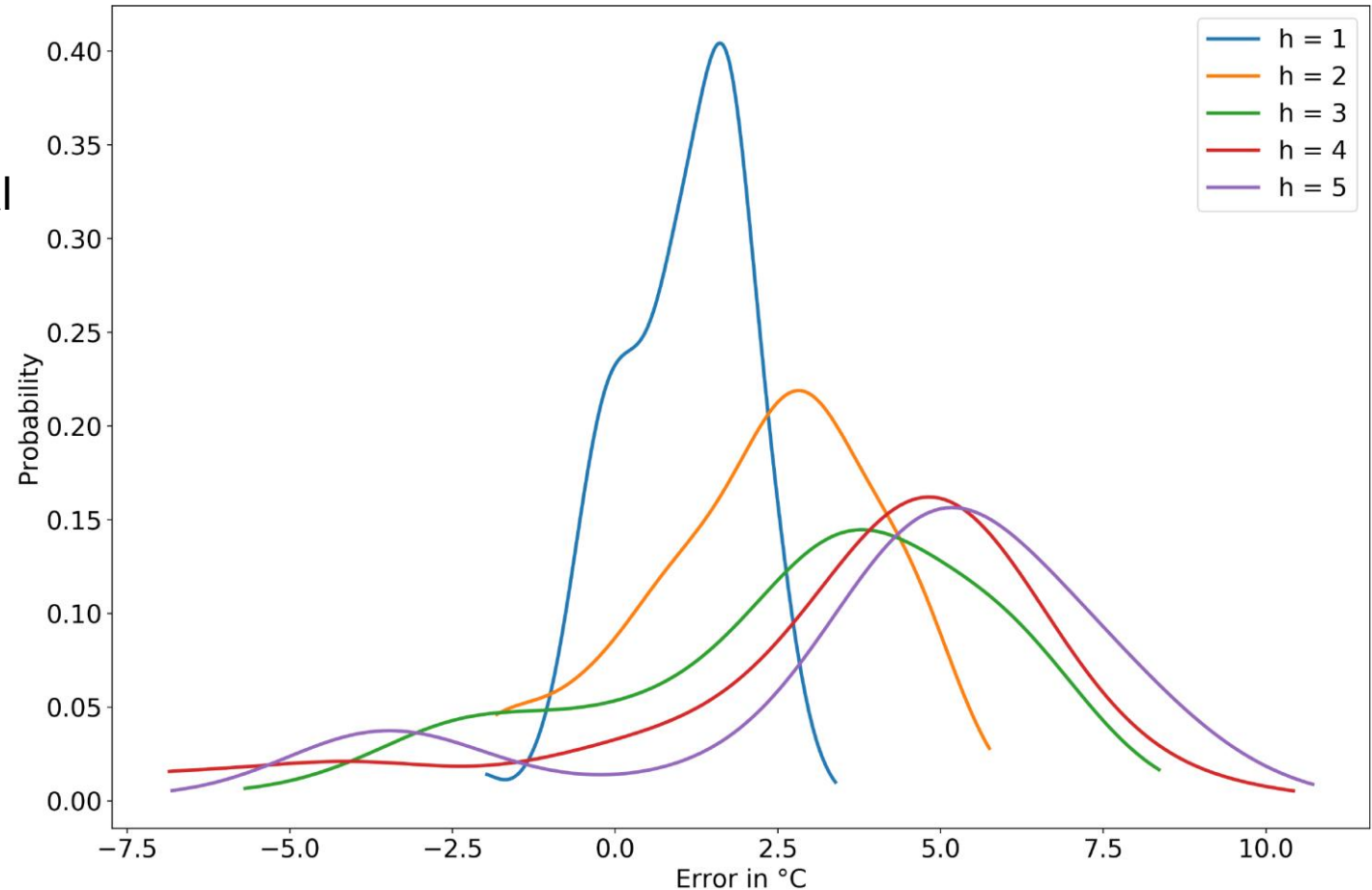
- Create a probabilistic forecast to project all past datapoints to the timestamp of the analysis.
- The probabilistic forecast assigns a distribution to all observations.
- Use shape of distribution to quantify the uncertainty.



# Probabilistic Forecast using Empirical Prediction Intervals

- Empirical prediction intervals [1] can be used to transform deterministic forecasts into probabilistic forecasts by adding the distribution of historical forecast errors to the deterministic forecast (dressing).
- Error distributions are location and forecast horizon dependent.

Example of error distributions for different forecast horizons



[1] Yun Shin Lee and Stefan Scholtes. Empirical prediction intervals revisited. *International Journal of Forecasting*, 30(2):217–234, 2014.

# Joint Uncertainty Estimation

- Required: Pair-wise **joint (un)certainty** of all observations which can be estimated via the joint distribution of two observations.
- Problem: Via the forecast we only have access to the (non-independent) marginal distributions of each observation. Estimating joint distribution from marginals is often a very complicated problem.
- Hint: Estimating (tight) bounds of the sum of two (potentially) dependent marginal distribution is much easier [1].

$$u_{ij} \leq P(\mathbf{X} + \mathbf{Y} \leq e)$$

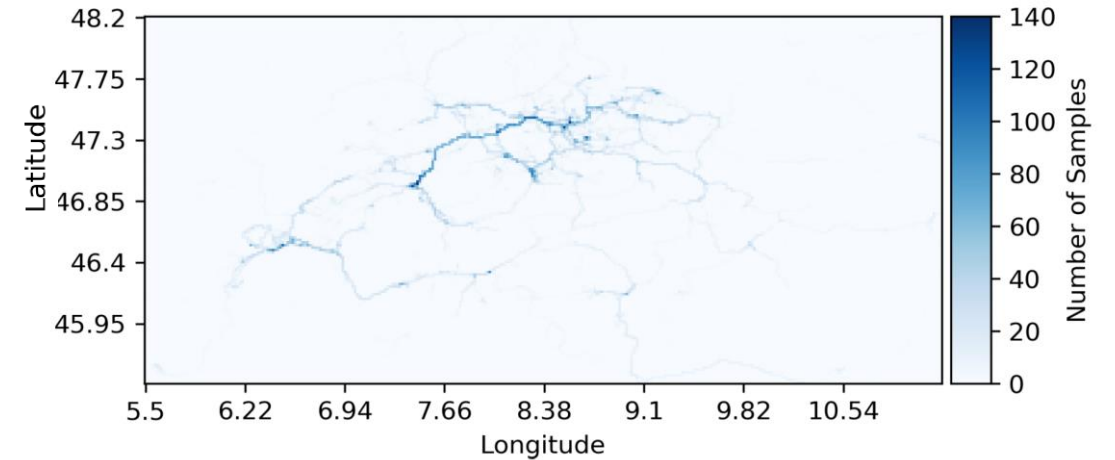
- $u_{ij}$  is the (lowest) probability that the sum of the marginals (absolute forecast errors) is below the threshold  $e$ . Therefore,  $u$  is a measurement of joint certainty.

[1] Michel Denuit, Christian Genest, and Étienne Marceau. Stochastic bounds on sums of dependent risks. *Insurance: Mathematics and Economics*, 25(1):85–104, 1999.

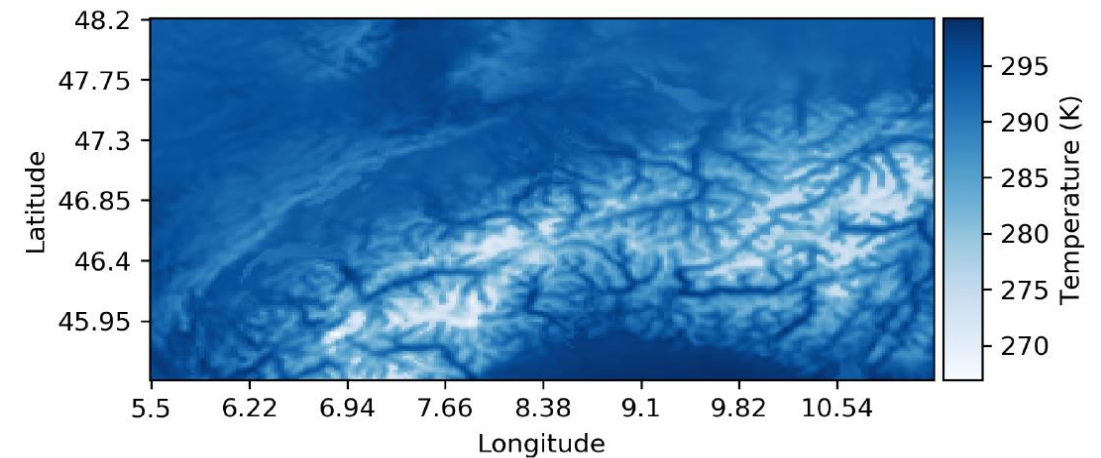
# Case Study 1 – Real data: Temperature sensors on cars

- GPS data of car movement from a tracking study in Switzerland are used as moving sensors.
- Car data is used to sample from a  $0.018^\circ$  cell grid (~2x2 km in Switzerland) with ambient temperature data (max 1 sample per cell).

Density of measurements (based on movement data)



Temperature in Switzerland (2nd of July 2013 09:00)





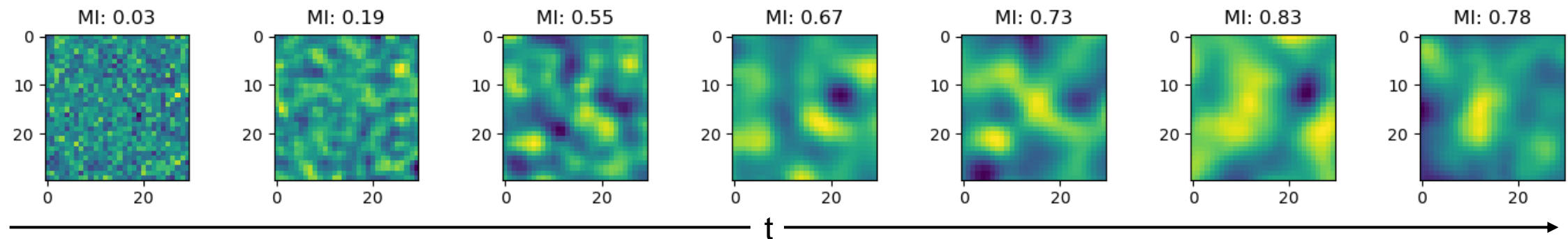
# Case Study 2 - Synthetic Data with Spatial Non-Stationarity

- Generate 60 time steps with a 60x60 grid cells resolution.
- Simple Kriging based on a Gaussian spatiotemporal variogram.
- Data has non-stationary spatial autocorrelation.

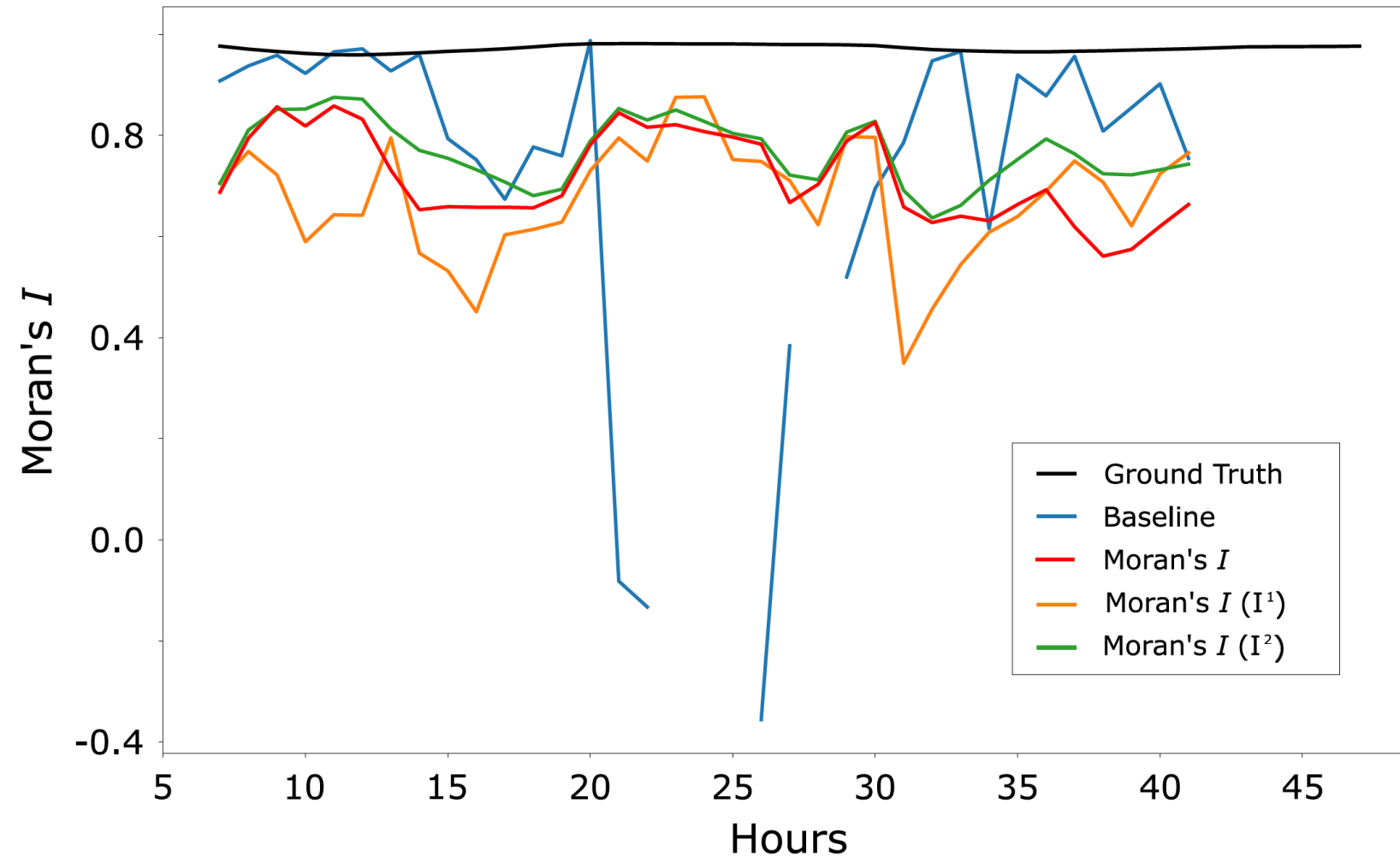
$$\gamma(h, r_i) = (s - n) \left( 1 - \exp \left( -\frac{h^2}{\frac{1}{3}r_i^2} \right) \right) + n$$

$r_i$ : range  
 $s$ : sill (1)  
 $n$ : nugget (0)

$$r_i = 0.5 + |10 \cdot \sin(2i/2\pi)|$$

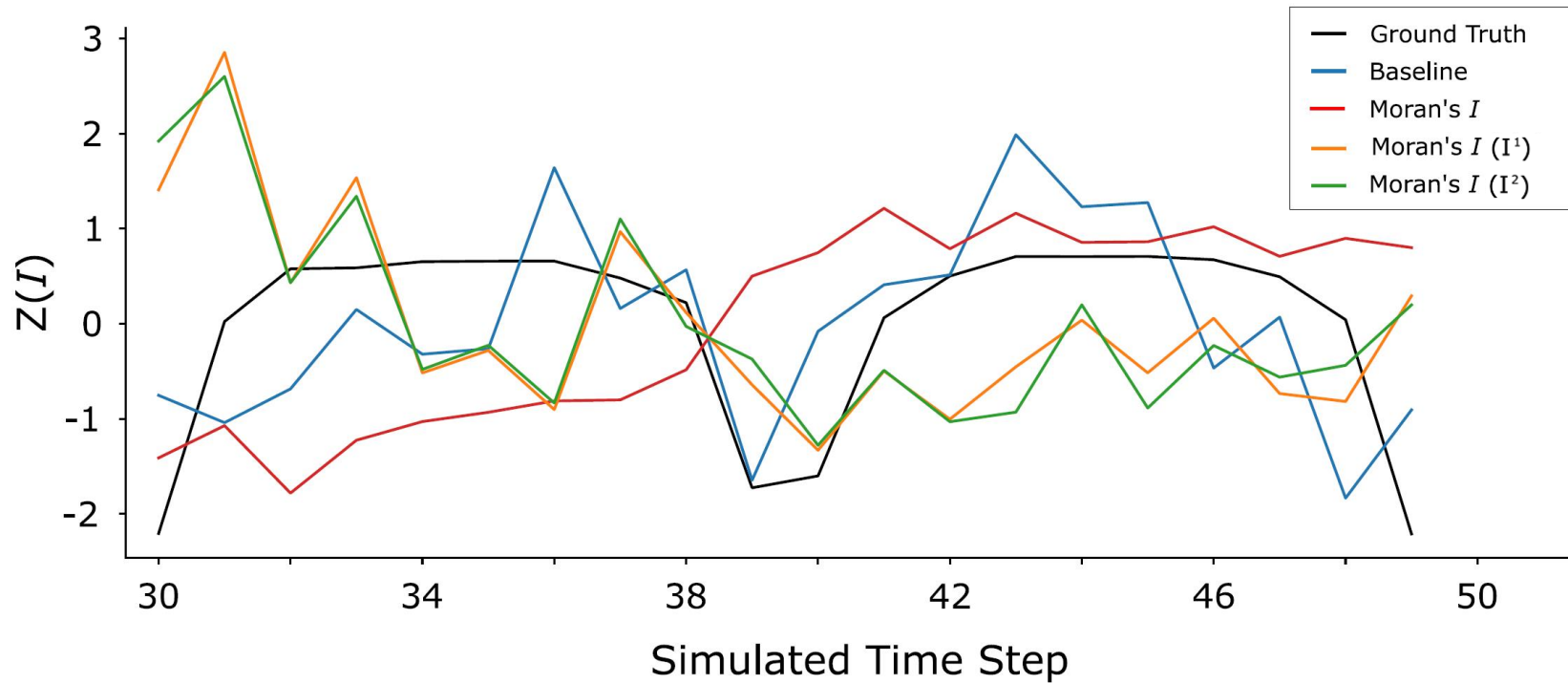
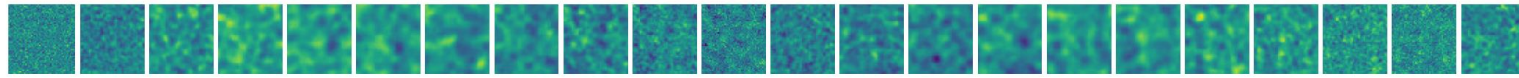


# Results Case Study 1



- Moran's  $I^2$  shows improvement over other methods.
- $I^1$  and  $I^2$  stable in periods with low data availability (causes gaps in baseline).
- As opposed to  $I^2$ ,  $I^1$  does underestimate Moran's  $I$  by design. Especially in a regime of low certainty.
- $I^2$  is less stable and more volatile than  $I^1$ .

# Results Case Study 2

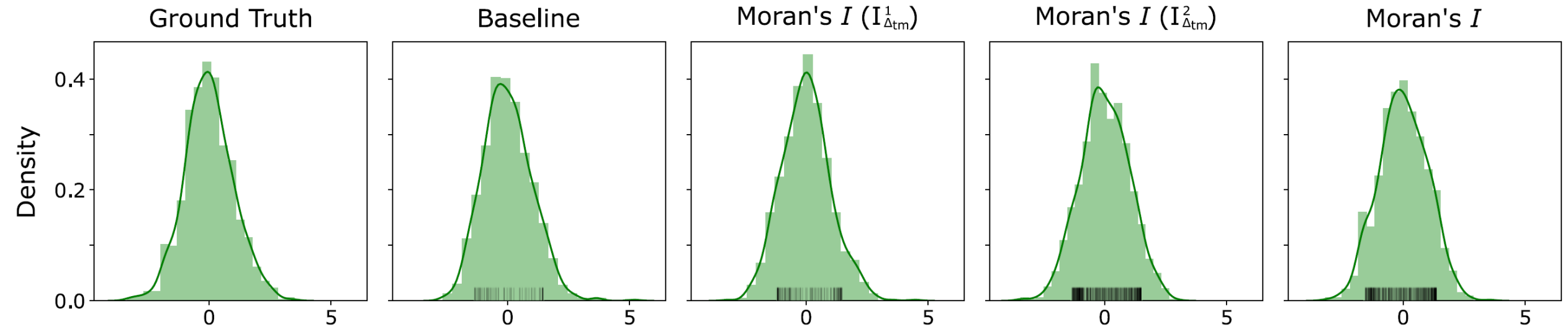


- $I^1$  and  $I^2$  not suitable for dealing with non-stationarity.
- Due to the low spatio-temporal correlation between time steps, baseline has best performance.

# Conclusions & Future Work

- Two novel ways to incorporate uncertainty related to moving sensors in the calculation of Moran's  $I$ .
- We show that incorporating uncertainty as a relative metric has significant advantages over the approach that incorporates absolute uncertainty values.
- Extend framework to local Moran's  $I$  to take into account spatial heterogeneity.
- Analyse impact of forecast performance on Moran's  $I$  estimate.

# Back Up



# Main Contributions

- Presentation of two modifications of Moran's I that take pair-wise spatio-temporal uncertainty of mobile sensor readings into account.
- Non-parametric approach for the estimation of pair-wise spatio-temporal uncertainty of mobile sensor readings.
- Validation of methods using two case studies.

# Probabilistic Forecast using Empirical Prediction Intervals

Empirical prediction intervals [1] can be used to transform deterministic forecasts into probabilistic forecast using historical forecast errors (dressing).

$$\hat{O}_{t_n, h} = f(O_{t \leq t_i})$$

$$e_{t_n, h} = O_{t_n} - \hat{O}_{t_n, h}$$

$$\hat{F}_h(e) = k^{-1} \sum_{t=1}^k \mathbb{I}(e_{t, h} \leq e)$$

$$\hat{O}_{t_n, h} + \hat{F}_h(e) = \text{probabilistic forecast}$$

[1] Yun Shin Lee and Stefan Scholtes. Empirical prediction intervals revisited. *International Journal of Forecasting*, 30(2):217–234, 2014.

$\hat{O}_{t_n, h}$ : Forecasted observation for time  $t_n$  with horizon  $h$

$e_{t_n, h}$ : Forecast error for time  $t_n$  with horizon  $h$

$\hat{F}_h(e)$ : Empirical cumulative distribution function of Forecast errors



# Uncertainty Estimation 2

- Required: Pair-wise **joint uncertainty** of all observations which can be estimated via the joint distribution of two observations.
- Problem: Via the forecast we only have access to the (non-independent) marginal distributions of each observation.
- Approach: Estimate (tight) lower bound of the probability that the forecast error (of the sum of both observations) is below a specified threshold [1].

$$u_{o_{il \rightarrow n}, o_{jm \rightarrow n}} = b_l(e) = \sup_{x \in \mathbb{R}} \max\{F_1^-(x) + F_2^-(e - x) - 1, 0\} \leq P(\mathbf{X} + \mathbf{Y} \leq e)$$

$u_{o_{il \rightarrow n}, o_{jm \rightarrow n}}$ : Joint certainty of  $o_{il}$  and  $o_{jm}$  projected to  $t_n$

$b_l(e)$ : Lower bound of probability to stay below forecast error

$e_{t_n h}$ : Forecast error for time  $t_n$  with horizon  $h$

$F_{1/2}^-$ : Left side limit of the CDF of the 1<sup>st</sup> or 2<sup>nd</sup> marginal distribution

$u_{o_{il \rightarrow n}, o_{jm \rightarrow n}}$  is the (lowest) probability that the sum of the absolute errors (of both forecasts) is below the threshold  $e$ . Therefore,  $u$  is a measurement of certainty.

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